

Low-Degree Outcomes and Clustered Designs: A Combined Approach for Causal Inference Under Interference



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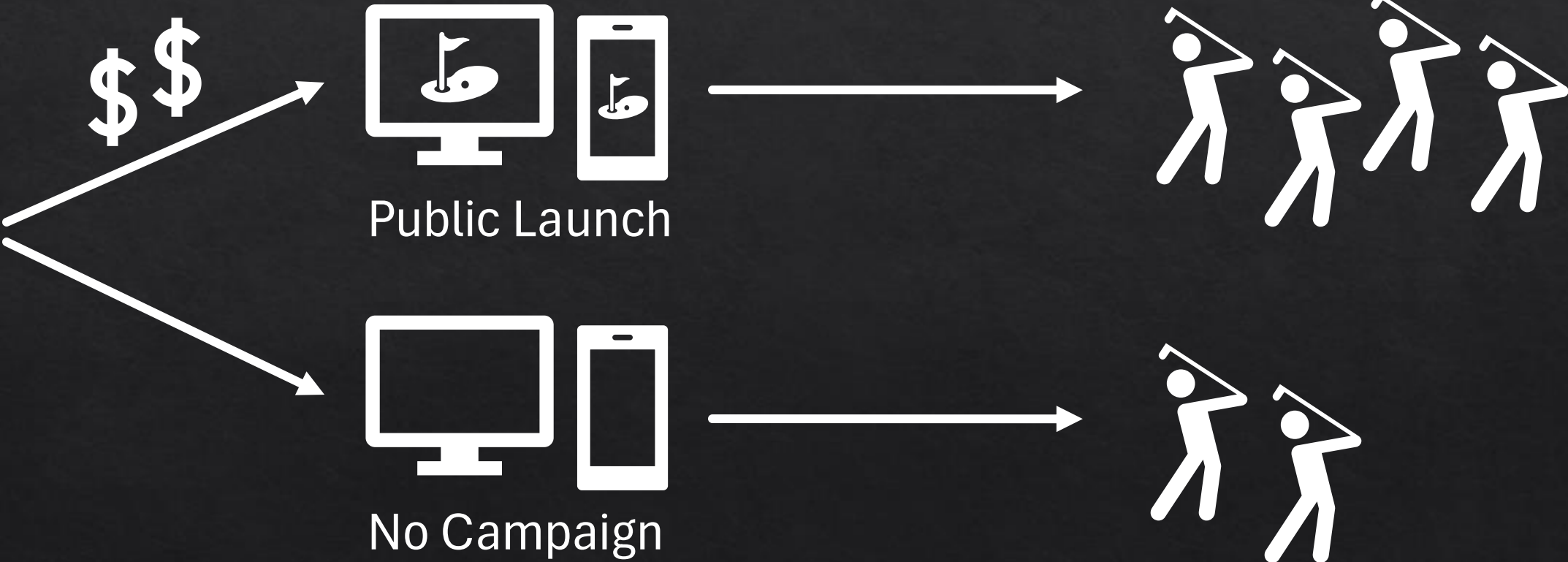
Johan Ugander



Christina Lee Yu

Motivation: Advertising

A golf course is deciding whether to run an advertising campaign



Motivation: Advertising

Use a (small) randomized experiment to inform their decision



Treatment Group



Control Group



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Treatment Group

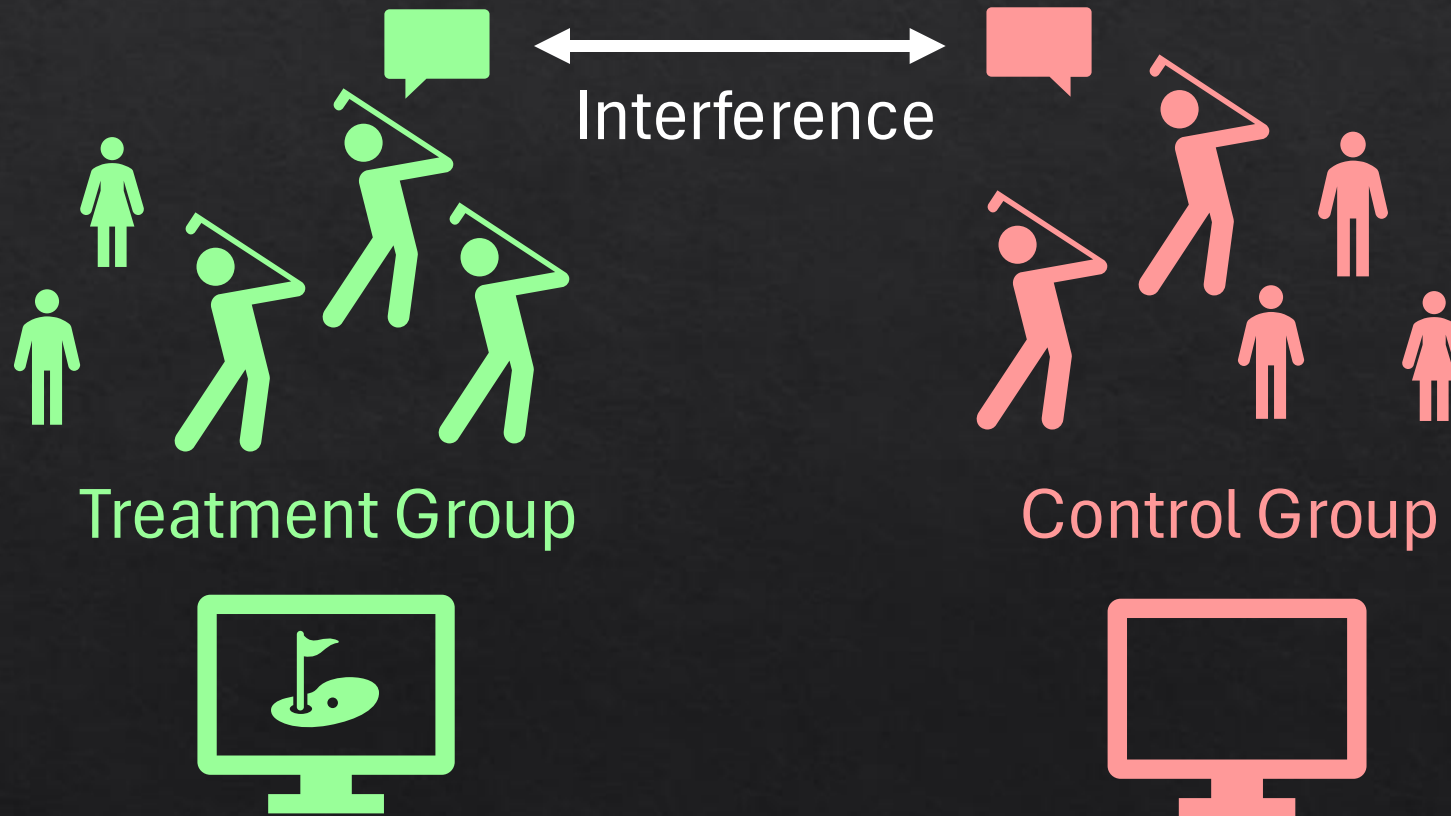


Control Group



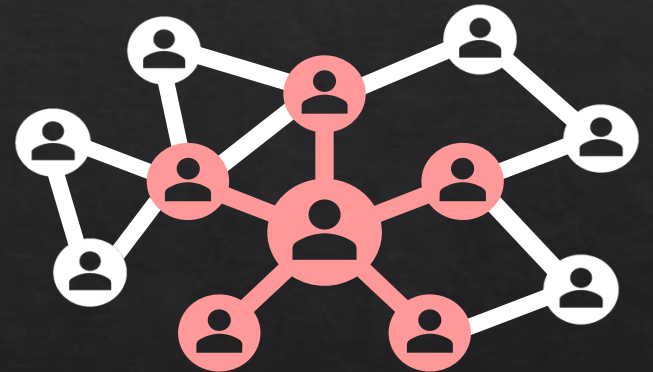
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Formalizing the Problem

- n individuals
- Treatment Assignments $z_i \in \{0,1\}$ (0 = control, 1 = treatment)
- Potential Outcomes $Y_i(\bar{z}) : \{0,1\}^n \rightarrow \mathbb{R}$
- Neighborhood Interference
- Goal: Estimate Total Treatment Effect



$$TTE = \frac{1}{n} \sum_{i \in [n]} (Y_i(\bar{\mathbf{1}}) - Y_i(\bar{\mathbf{0}}))$$

A Simple Unbiased Estimator

- Independent Treatment Assignments $z_j \sim \text{Bernoulli}(p)$
- Horvitz-Thompson Estimator

$$\widehat{TTE}_{\text{HT}} = \frac{1}{n} \sum_{i \in [n]} Y_i(\bar{\mathbf{z}}) \left(\underbrace{\prod_{j \in N_i} \frac{z_j}{p}}_{\substack{0 \text{ unless entire} \\ \text{neighborhood} \\ \text{treated}}} - \underbrace{\prod_{j \in N_i} \frac{1-z_j}{1-p}}_{\substack{0 \text{ unless entire} \\ \text{neighborhood} \\ \text{untreated}}} \right)$$

Prohibitively High Variance $O\left(\frac{1}{p^d}\right)$

Two Approaches for Variance Reduction

1. Better Experimental Design

- Under Bernoulli treatment, most neighborhoods are partially treated
- Smarter designs (e.g., with clustering) increases prevalence of fully treated neighborhoods
- Variance reduction relies on structural assumptions of **causal network**

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- More clever estimators can utilize measurements from partially treated neighborhoods
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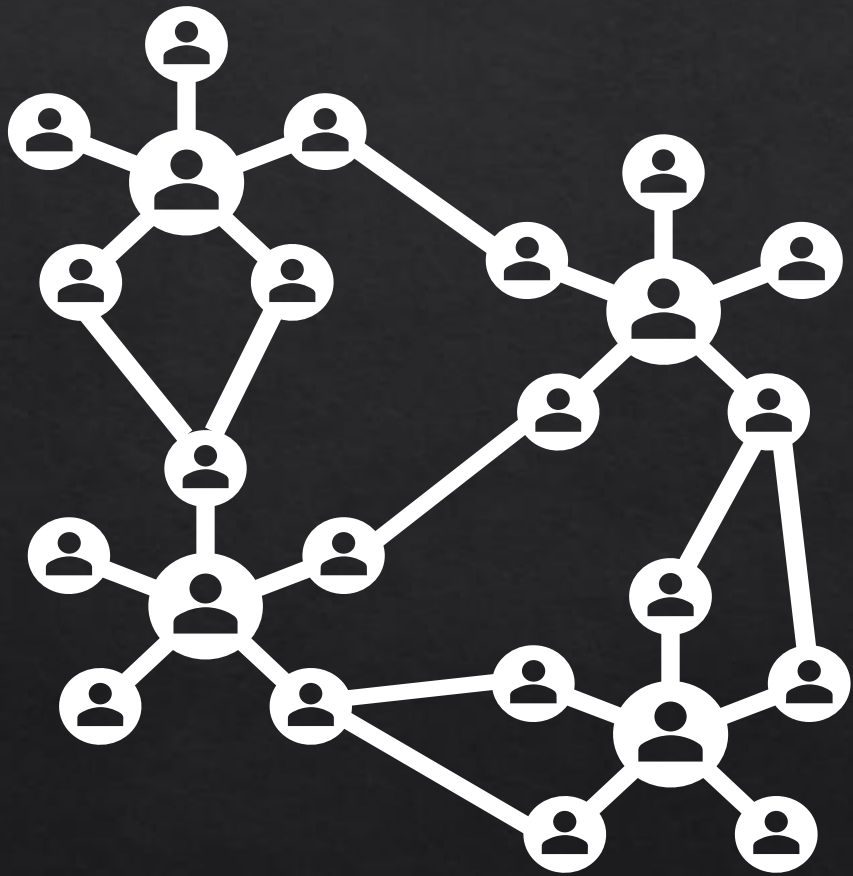
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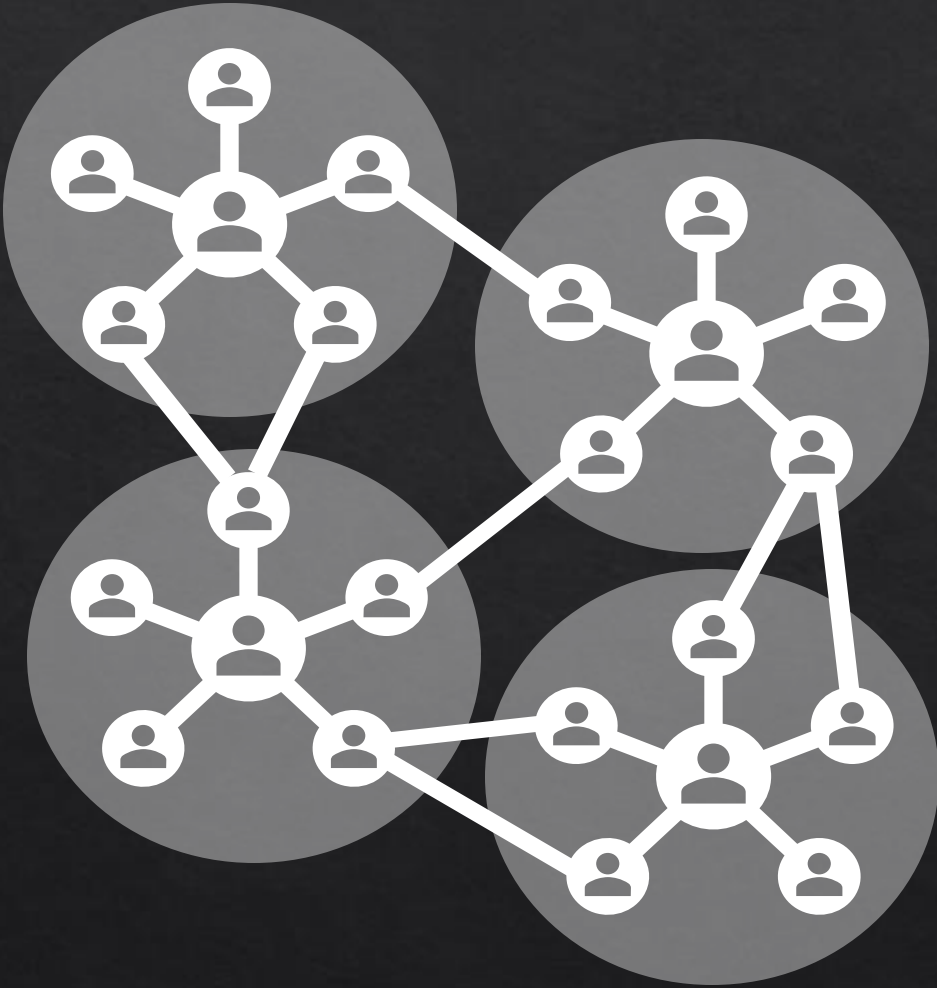
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Why not both?

Graph Cluster Randomized Design

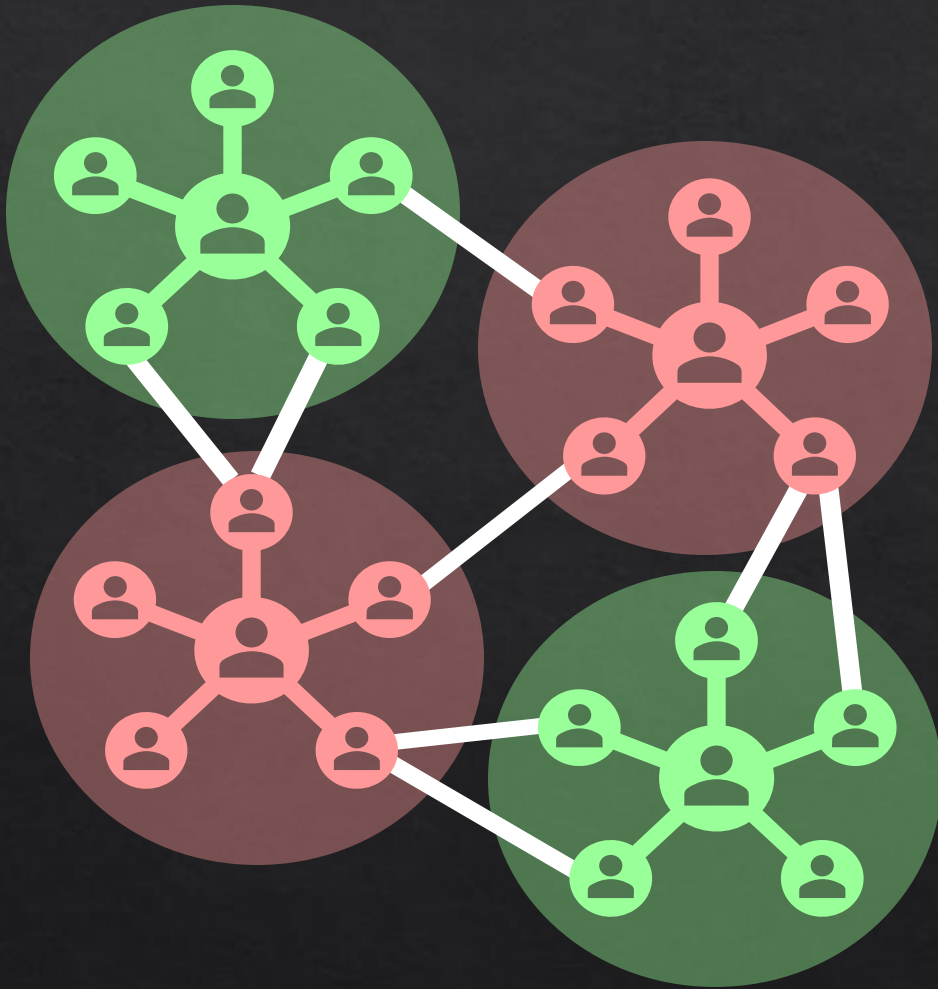


Graph Cluster Randomized Design



- Partition a graph into k clusters
 - $\mathcal{C} = \{C_1, C_2, \dots, C_k\}$
 - $\mathcal{C}(i) =$ cluster of individual i
 - $\mathcal{C}(N_i) = \bigcup_{j \in N_i} \mathcal{C}(j)$

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 - $\mathcal{C}(i) = \text{cluster of individual } i$
 - $\mathcal{C}(N_i) = \bigcup_{j \in N_i} \mathcal{C}(j)$
- Sample cluster treatments
 - $w_{C_\ell} = \mathbb{I}(C_\ell \text{ treated}) \sim \text{Bernoulli}(p)$
- Assign individual treatments
 - $x_i = w_{\mathcal{C}(i)}$

Variance Bounds for GCR

- For Bernoulli (unit randomized) Design:

$$\text{Var}(\widehat{TE}_{HT}) = O(n^{-1}p^{-d})$$

Network Degree

- For Graph Cluster Randomization:

$$\text{Var}(\widehat{TE}_{HT}) = O\left(n^{-1}p^{-\max_i |c(N_i)|}\right)$$

“Cluster Degree”

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- 3-Net Clustering on κ -restricted growth graphs

$$\text{Var}(\widehat{TE}_{HT}) = O(n^{-1}d \kappa^5 p^{-\kappa^6})$$

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- Incorporate information from partially-treated neighborhoods
- Need assumptions on potential outcomes



Alternate Estimator Designs

- Incorporate information from partially-treated neighborhoods
- Need assumptions on potential outcomes
- Low-order Interactions Model



$$Y_i(\bar{\mathbf{z}}) = \sum_{S \in \mathcal{S}_i^\beta} c_{i,S} \prod_{j \in S} z_j = \langle \mathbf{c}_i, \tilde{\mathbf{z}}_i \rangle \quad \mathcal{S}_i^\beta = \text{subsets of } N_i \text{ with size } \leq \beta$$

The Pseudoinverse Estimator

- Under Low-order Interactions, we estimate $\hat{c}_i = Y_i(\bar{\mathbf{z}}) \mathbb{E}[\tilde{\mathbf{z}}_i \tilde{\mathbf{z}}_i^T]^\dagger \tilde{\mathbf{z}}_i$

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- Extend by linearity to estimate $\widehat{TTE}_{PI} = \frac{1}{n} \sum_{i \in [n]} \langle \hat{c}_i, \theta_i \rangle \quad \theta_i = (0, 1, \dots, 1)$

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- Extend by linearity to estimate $\widehat{TTE}_{PI} = \frac{1}{n} \sum_{i \in [n]} \langle \hat{c}_i, \theta_i \rangle \quad \theta_i = (0, 1, \dots, 1)$
- For Bernoulli (unit randomized) Design:

$$\widehat{TTE}_{PI} = \frac{1}{n} \sum_{i \in [n]} Y_i(\bar{z}) \sum_{S \in S_i^\beta} \left(\frac{1}{p^{|S|}} - \frac{1}{(1-p)^{|S|}} \right) \prod_{j \in S} (z_j - p)$$

- Unbiased with $\text{Var}(\widehat{TTE}_{PI}) = O\left(\frac{d^2}{n} \left(\frac{ed}{\beta p^2}\right)^\beta\right)$

Combining these Approaches

- PI estimator works for any experimental design:

Main Theorem:

\widehat{TTE}_{PI} is unbiased when each θ_i lies in the column space of $\mathbb{E}[\tilde{\mathbf{z}}_i \tilde{\mathbf{z}}_i^T]^\dagger$ with:

$$\text{Var}(\widehat{TTE}_{PI}) = O\left(\frac{1}{n^2} \sum_{i,j \in [n]} \gamma_i \gamma_j \mathbb{I}(\tilde{\mathbf{z}}_i \not\perp \tilde{\mathbf{z}}_j)\right)$$

where $\gamma_i = \sqrt{|s_i^\beta| \theta_i^T \mathbb{E}[\tilde{\mathbf{z}}_i \tilde{\mathbf{z}}_i^T]^\dagger \theta_i}$ is i 's contribution to the variance

and $\mathbb{I}(\tilde{\mathbf{z}}_i \not\perp \tilde{\mathbf{z}}_j)$ captures the graph structure's effect on the variance

GCR + PI Estimator

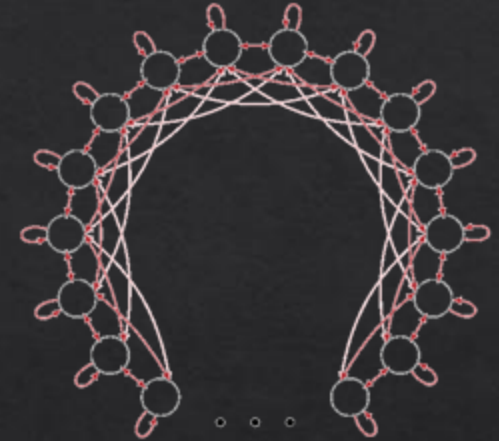
$$\gamma_i^2 \leq \begin{cases} 2d_i^\beta \cdot p^{-|\mathcal{C}(N_i)|} & |\mathcal{C}(N_i)| < \beta \\ 2d_i^\beta \cdot |\mathcal{C}(N_i)|^\beta \cdot p^{-\beta} & |\mathcal{C}(N_i)| \geq \beta \end{cases}$$

- When clustering quality “better than” potential outcomes complexity, variance depends exponentially on clustering quality
- When potential outcomes complexity is “better than” clustering quality, variance depends exponentially on interaction size parameter β

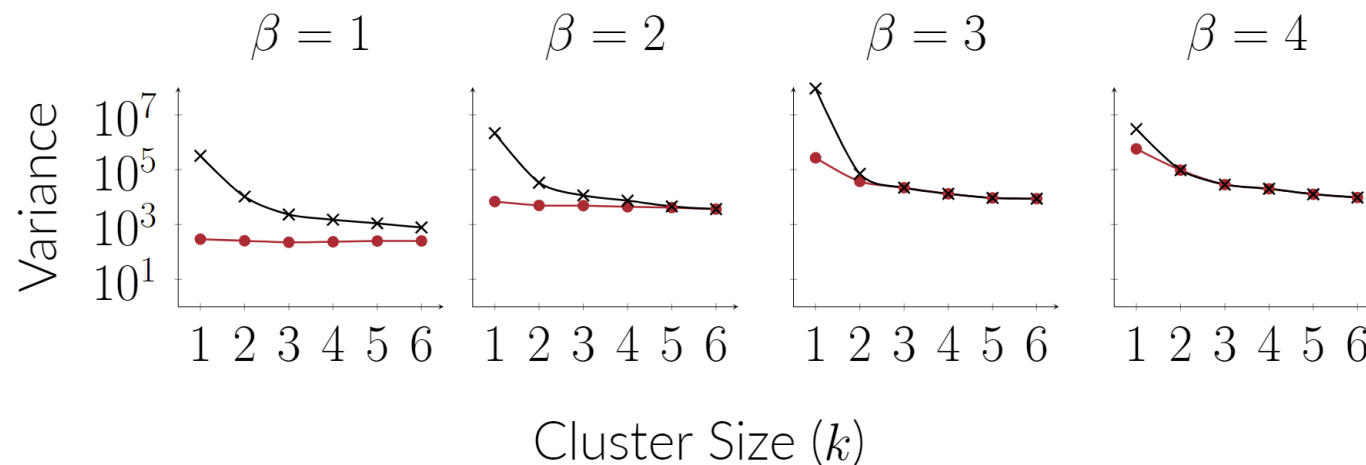
“Best of Both Worlds” Trade-off!

Experimental Results

- Cycle network with $n = 120$ and degree $d = 7$
- Effect coefficients sampled independently with $c_{i,S} \sim \text{Unif}(0, 10 \cdot 2^{-|S|})$
- Treatment probability $p = 0.25$
- Cluster vertices into adjacent groups
- Vary cluster size k and interaction size parameter β



Horvitz-Thompson (HT) vs. Pseudoinverse (PI)



Thanks for Listening!